

The Forward March of Categorical Tolerance in the United States

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Abstract: This article updates the empirical picture of categorical tolerance (CT), namely, the pattern of refusing to report dislikes across cultural genres, for the third decade of the twenty-first century in the United States. Analyzing recent survey data from two platforms, I find that CT has continued its march among Americans, reaching approximately one in five respondents. The analysis confirms earlier-observed demographic trends, showing that CT is strongly associated with younger cohorts and non-white individuals. However, I also find that individuals reporting the highest educational attainment are now overrepresented among categorical tolerants, suggesting that CT may increasingly function as an elite cultural strategy consistent with contemporary forms of status display, signaling openness and refusal to refuse. Furthermore, I find that while the odds of being a CT are not strongly polarized by political ideology, the inclination toward symbolic exclusion among non-CTs is, with conservatives significantly more likely to express a greater volume of cultural dislikes than liberals.

Keywords: categorical tolerance; cultural taste; symbolic exclusion; political ideology

Reproducibility Package: Data files and *R* code (in Quarto Markdown format) necessary to reproduce all of the analyses, tables, and figures reported in the article can be found at the following GitHub repo: <https://github.com/olizardo/sociological-science-categorical-tolerance-followup>.

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EVER since Richard Peterson introduced the felicitous term “patterns of cultural choice” into sociological discourse more than four decades ago (Peterson 1983), sociologists have been busy empirically uncovering and debating the sociological reality, drivers, meaning, and correlates of various such patterns. These include, most famously and influentially, the Petersonian cultural *omnivore* (Goldberg 2011; Peterson and Kern 1996; Warde, Wright, and Gayo-Cal 2008) along with their evil twin the culturally exclusive *univore* (Bryson 1997; Peterson 1992; Tonković et al. 2025), and also culturally *voracious* consumers (Sullivan and Katz-Gerro 2007), *poly-purists* and *poly-mixers* (Goldberg, Hannan, and Kovács 2016; Poecke 2018), *decisives* and *indecisives* (Sonnett 2004), and even plain old *snobs* engaging in (presumably) Bourdieusian boundary-drawing distinction (Atkinson 2011; Nault et al. 2021), among many others.

In a article published almost a decade ago, Lizardo and Skiles (2016b)—hereafter L&S—introduced another such pattern into sociological discourse, which they labeled *categorical tolerance* (CT), to refer to the phenomenon of a growing proportion of the population who refuse to report dislikes of *any* cultural genres—typically musical genres—in standard social science surveys designed to measure taste and consumption practices. In that article, L&S noted that CT appeared to be on the rise in the United States, going from about 5 percent of the population in the

canonical 1993 General Social Survey culture module data set to about 16 percent in a then-recent survey of a representative sample of Americans collected in 2012, which matched the GSS 1993 battery of items on likes and dislikes of musical tastes across 18 genres. This resulted in a significant shift in the shape of the univariate count distributions of number of dislikes, which went from a standard Poisson distribution in the GSS 1993 to a count distribution with excess zeroes in 2012 (Lizardo and Skiles 2016b: p. 90, Figure 1).

In the same article, L&S also showed that the rise of CT among Americans was a socially patterned phenomenon. Particularly, L&S showed that refusal to dislike was a growing trend among more recently born cohorts, suggestive of a macro-level process of cultural change via cohort replacement of the type that has been shown to govern macro-level cultural change across a wide variety of values, attitudes, beliefs, and practices (Kiley and Vaisey 2020; Ochoa and Vaisey 2024; Vaisey and Lizardo 2016; Vaisey and Kiley 2021). Moreover, when comparing 2012 to 1993 data, L&S found that categorical tolerants were over-represented among non-white respondents, with a slight under-representation of people with a college degree, indicating that CT did not seem to be driven by the usual status and cultural-capital-based dynamics emphasized by Bourdieusian models of social distinction (Bourdieu 1984), which usually focuses on status display among higher educated elites (Jarness and Friedman 2017; Khan 2011).

This article uses recently collected data from a representative sample of the American population to update the empirical picture of CT in the United States for the third decade of the twenty-first century. First, I ask whether we can observe a continued rise in the CT effect in the more recent data. Second, I examine the socio-demographic correlates of CT to determine whether this pattern of cultural choice persists in a similar socio-demographic “niche” (McPherson 2004) as L&S observed in 2012, or whether its location in social space has evolved since then. Third, I also examine the socio-demographic correlates of *symbolic exclusion* among people at risk of expressing dislikes (Bryson 1996), to see if we can observe similar patterns to those previously uncovered by L&S. Finally, the article adds a missing dimension to previous work on CT and symbolic exclusion by examining whether these cultural taste patterns are connected to political identification as liberal or conservative, given recent work indicating that cultural taste patterns have themselves become embroiled in the “oil spill” (DellaPosta 2020) produced by the continuing political polarization of the American mass public in the past two decades, a phenomenon that has accelerated in the Trump era (Rawlings and Childress 2023).

To preview the main results: I find that CT has indeed continued its onward march among Americans, with about 20 percent of respondents failing to report disliking any of the 12 genres in data collected from three survey samples on two separate platforms. Second, the social location of CT is fairly similar to that observed by L&S a decade ago; younger adults are overrepresented among categorical tolerants relative to older adults. Similarly, as L&S found in 2012, non-white respondents are overrepresented among categorical tolerants in the United States. One distinct difference between the current findings and L&S’s results, however, is a seeming trickle-up of CT as an elite taste pattern. Those reporting the highest levels of education are overrepresented in this group, indicating that Bourdieusian status display dynamics may increasingly be at play. Finally, while CT is not polarized

along liberal/conservative lines, *symbolic exclusion* patterns net of CT are. After adjusting for other socio-demographic predictors, conservatives are more likely to express a greater number of cultural dislikes than liberals. This indicates that political ideology, along with age—but not education—is now the most reliable predictor of expressing dislikes for cultural genres in the United States today.

The rest of the article is organized as follows. The next section briefly reviews recent work that builds on L&S's idea of CT, as well as other work examining the growing linkages between cultural tastes and political ideology. The next sections introduce the data and outline the basic analytic plan, presenting the main results. Finally, the discussion section summarizes the key findings, highlights their main implications, and suggests avenues for future work.

Recent Work on Taste and CT

The traditional view that the upper classes use “highbrow” cultural tastes to symbolically exclude others and reproduce social hierarchies—uncritically associated with the work of (Bourdieu 1984)—is typically contrasted to the “cultural omnivore” thesis, first articulated by Peterson (1992). This influential proposal suggests that the patterns of cultural choice of those at the top of the status ladder (typically indexed by high levels of educational attainment) are increasingly characterized by an appreciation and enjoyment of a broad range of cultural activities, spanning both traditionally “highbrow” (e.g., opera, classical music, and art museums) and “lowbrow” genres (e.g., heavy metal and sporting events) (Goldberg 2011; Lizardo and Skiles 2012; Peterson and Kern 1996). This shift in the social bases of distinction from snobbery to omnivorousness blurs old boundaries of cultural hierarchy, interpreted as a sign that “conspicuous” tolerance and openness to cultural diversity may be the new form of high-status culture (Ollivier 2008).

The general empirical implication is that cultural omnivores should be more (explicitly) tolerant, liberal, and inclusive, displaying a fundamental openness to consume and engage in a wide variety of cultural practices and experiences. However, the Petersonian thesis does not imply that omnivores indiscriminately accept all cultural forms. Instead, subsequent research goes to some lengths to emphasize that cultural omnivores still engage in “patterned tolerance” and the selective exclusion of low-status cultural forms (Bryson 1996; Lindblom and Mustonen 2019; Prieur, Rosenlund, and Skjøtt-Larsen 2008; Warde et al. 2008). This means that while cultural elites embrace a wide variety of cultural genres, they often express explicit dislikes toward and symbolic distancing from a few specific genres, particularly those associated with lower-status groups (Bryson 1996; Lindblom 2022).¹

Why are dislikes expressed in social surveys on cultural tastes important? As Bourdieu (1984: 56) himself noted, taste is often asserted negatively via refusal. Bryson (1996) linked this dynamic to the idea of “symbolic exclusion” or the *indirect* action of drawing a symbolic (social) boundary against particular (usually lower status) groups by distancing oneself from the cultural objects that those groups are perceived to consume (Lizardo and Skiles 2016a). This recognition has led scholars to stress the critical importance of studying “dislikes” alongside “likes” to fully understand the dynamics of symbolic exclusion and the construction of cultural boundaries (Lindblom 2022; Lizardo and Skiles 2015; Oncini et al. 2023).

Importantly, this general framework has proven helpful for understanding social and cultural dynamics, beyond the case of the United States. For example, a recent study of musical taste conducted in South Korea found that tolerant omnivores still excluded genres such as “gugak” (Korean traditional music) and “trot,” (one of the oldest forms of Korean popular music), which are preferred by lower-educated and older individuals, suggesting that the more “cosmopolitan” omnivores still engage in patterns of symbolic exclusion directed at older persons of lower educational status (Kim et al. 2020: 321).

Cultural Tolerance and Social Tolerance

As noted, since the seminal work of Bryson (1996), the interplay between cultural taste orientations and broader social attitudes has been an active area of research. Bryson’s original study showed that cultural dislikes predicted traditional socio-political attitudes associated with political tolerance and symbolic racism; people who engaged in more symbolic distancing in the realm of culture (via expressing a larger number of dislikes) were also more likely to have higher scores in the General Social Survey’s symbolic racism and political intolerance scales.

Various studies conducted since Bryson’s consistently show an empirical link between cultural openness and social tolerance across multiple contexts. For instance, Cho and Kim (2023) show that cultural omnivores in South Korea are more accepting of immigrants, demonstrating a greater willingness to include them as friends, neighbors, and even marriage partners. Beyond immigration, cultural omnivorousness is positively correlated with a liberal political outlook and a higher degree of social trust, particularly toward outgroup members. Cultural omnivorousness is also associated with cosmopolitan values, support for social integration and inclusiveness, and progressive stances such as the rejection of traditional gender roles and acceptance of alternative sexual lifestyles (Chan 2019). In line with this work, a study conducted in Finland by Lindblom (2022) indicates that omnivores are more likely to hold liberal views on immigration, law enforcement, and same-sex relationships, while also prioritizing environmental protection (Lindblom 2022).

Conversely, Univorism, a concept first conceptualized by Peterson (1992) and empirically studied by Bryson (1997), describes individuals characterized by narrow, limited cultural tastes, typically confining themselves to a more delimited range of popular cultural practices (Lindblom 2022; Cho and Kim 2023). This restricted taste profile is often characteristic of older age groups and those with lower educational attainment (Lindblom and Mustonen 2019; Sokolova and Sokolov 2020). Recent work shows that univores exhibit less liberal and less tolerant attitudes across various social and political domains. For instance, univores tend to be positioned on the conservative end of the liberal-conservative spectrum, are more skeptical toward fellow citizens, indicating lower social trust, have lower voter turnout, are more critical of immigration, and generally do not prioritize environmental protection (Chan 2019; Cho and Kim 2023; Lindblom 2022).

The Rise of CT

The concept of CT, as introduced by L&S, identifies a cultural taste orientation characterized by an individual’s refusal to express dislikes toward any cultural

genre or item. In this way, CT effectively translates into the refusal to show explicit forms of symbolic exclusion (Bryson 1996). As noted, a lower likelihood to engage in symbolic exclusion via cultural dislikes is a taste pattern traditionally linked to individuals of high social status, implying a form of “openness to cultural diversity” that presumably serves as a new ethos for privileged strata (Lindblom and Mustonen 2019; Ollivier 2008). Conceptually, however, CT can be positioned as a distinct taste orientation relative to cultural omnivorousness, which involves a wide range of “likes” and very few (but not the absence of) “dislikes,” and univorousness, characterized by a narrow range of “likes” and numerous “dislikes” (Sonnett 2004). This orientation suggests a potential shift in how individuals navigate cultural boundaries and express particular forms of taste in the current setting (Lizardo and Skiles 2016b).

In their original research using U.S. data, L&S found that the phenomenon of CT has notably increased over time among non-white and non-elite populations. Recent work by Lindblom (2022) most closely follows L&S’s original design, namely, comparing two repeated cross-sections of the same population. Using a survey instrument querying respondents on their likes and dislikes for 13 musical genres, Lindblom (2022) finds that the proportion of categorical tolerants (persons who refuse to dislike any of the genres) increased in Finland from about 14 percent of the population in 2007 to 20 percent in 2018. Regarding the social distribution of CT, aside from women being over-represented, Lindblom (2022) found few socio-demographic correlates but observed that CTs displayed greater ambivalence in their taste preferences (e.g., “mixed feelings”) compared to both omnivores and univores, with social non-distinctiveness also increasing over time (Lindblom 2022).

This suggests that for CTs, the absence of dislikes might stem from indecisiveness or a reluctance to take a definitive stance rather than from a genuine embrace of distinct cultural forms (Lindblom 2022). In the end, Lindblom (2022: p. 13) concluded that the taste orientation signaled by CT, while on the rise, perhaps signaled a kind of socially unstructured “ambivalence” (Sonnett 2004). However, unlike L&S’s study, Lindblom (2022) did not use a statistical approach (e.g., zero-inflated count models) to distinguish respondents at risk of expressing dislikes from true refusers. As such, whether CT has a coherent social niche in the contemporary context remains somewhat of an open question.

Subsequent work has extended the CT framework to study other forms of cultural taste beyond music. For instance, recent work identifies a segment of young urban Finns who exhibit CT in culinary tastes. Although these individuals exhibited fewer overall dislikes, individuals with higher educational credentials still demonstrated “patterned tolerance” (Bryson 1996), by selectively disliking certain “fast and convenient” cuisines, reinforcing class-based symbolic boundaries (Lindblom and Mustonen 2019). Conversely, among less educated groups, CT in food preferences was sometimes interpreted as a form of “cultural goodwill” (Bourdieu 1984), or an emulation of socially valued tastes on the part of middle-status persons (Lindblom and Mustonen 2019). In addition, in a study of food and drink avoidances in Italy (Oncini et al. 2023), a CT cluster was also identified (e.g., expressing little to no culinary dislikes), whose members avoided fewer items than average, resembled cultural omnivores, and were more likely to be younger, male, and possessing higher cultural and economic capital. Overall, these findings

highlight that the manifestation and social correlates of CT can vary, reflecting complex interactions between individual dispositions, socioeconomic factors, and specific cultural contexts (Lindblom and Mustonen 2019; Oncini et al. 2023).

To summarize, the recent literature on cultural taste has converged around two main themes. One strand of work, building on the foundational work of (Bryson 1996), has explored the connection, perhaps rising in recent times, between “cultural tolerance” (as expressed in the refusal to dislike other cultural forms) and various forms of attitudinal tolerance and “liberalism” across various socio-political, practical, and attitudinal domains. This work suggests that in the context of increasing political polarization and even rising political “sectarianism” (Finkel et al. 2020), we should observe a linkage between cultural taste and political ideology (Rawlings and Childress 2023), with liberals being less likely to engage in symbolic exclusion than conservatives.

Second, work on “CT” suggests that this particular pattern of taste, characterized by the universal refusal to engage in *any* form of symbolic exclusion may be on the rise, and could become characteristic of a particular (younger, less white) segment of the population in the United States; alternatively, as it has spread, CT may have lost its social moorings, not being strongly correlated with any markers of social position (Lindblom 2022). By the same token, if CT is part of the increasingly observed package of “conspicuous openness to diversity” evident in the avowed tastes of contemporary elites (Ollivier 2008), then it may have even undergone an upgrading in the social space, being more likely to be espoused by individuals with high cultural capital. The empirical analysis that follows addresses both themes in the context of the third decade of the twenty-first century in the United States.

Data

The data for this study come from three rounds of survey data collection on Americans’ cultural tastes. The first two waves of data collection yielded two convenience samples: one collected in 2021 and the other in 2023, the latter with an oversample of people with advanced educational degrees. Both of these samples come from the Prolific Academic (ProA) platform. Prolific has emerged as a preferred alternative to Amazon’s Mechanical Turk for researchers seeking high-quality, suitably motivated participant pools. Unlike MTurk, ProA was explicitly designed for social-scientific researchers and offers a high-quality, better-compensated participant pool. Also, unlike MTurk, ProA mandates a minimum payment based on the expected time required to complete the task. Participants in our study were compensated at a rate of \$9.50 per hour, which ProA indicates as a “good” compensation rate. The main participant pool for this study on the Prolific platform was limited to participants between the ages of 18 and 85 who reside in the United States and whose first language is English (or are fluent in English). The study was also restricted to participants who scored at or above 95 percent on ProA’s “approval rate” (the proportion of times a participant’s study response has been accepted in other studies as adequate). Survey completion was verified by embedding a completion code as the last question, which participants had to input on the Prolific website to confirm they had read and completed each survey question.

In the summer of 2025, a third sample was collected to replicate the results of the first two waves using a more representative sample of American respondents, using the relatively recently launched Cint (previously Lucid) Theorem platform, which provides access to a quota-sampled set of respondents matched to the U.S. census by age, gender, ethnicity, and region. Like Prolific, Cint Theorem is designed primarily for social science research. Cint employs a “marketplace” approach to recruiting study participants by connecting researchers with various suppliers and offering a variety of incentives to respondents. In this way, Cint provides relatively affordable access to participants for researchers (about \$1.50 per completed survey) and high-quality samples (Crabtree 2025). Cint also uses various automated screens (e.g., invalid traffic and duplicate IPs) and response quality indicators (e.g., CAPTCHA scores) to weed out bots and prevent less attentive participants from accessing the survey. Representative samples of the U.S. population from Cint Theorem have been used effectively in recent social scientific work on Americans’ political and social attitudes (e.g., Keskintürk, Kiley, and Vaisey 2025).

After eliminating completed surveys with an automated CAPTCHA score below the midpoint of 0.5 (and thus increased likelihood of being bots), we ended up with an $N = 555$ for the 2021 ProA sample, an $N = 853$ for the summer of 2023 ProA sample, and an $N = 959$ for the summer of 2025 Cint sample. The pooled sample across all three surveys is $N = 2367$.

Results

Probability of Dislikes by Genre

Table 1 shows the probability of any genre experiencing a judgment of dislike for both the Prolific and Lucid samples, pooling the 2021 and 2023 Prolific samples. For each genre, participants were asked to “Tell us your feelings about each type of music.” Evaluation options were anchored at the extremes with “dislike a great deal” and “like a great deal”; “neither like nor dislike” served as the neutral point. A participant was counted as expressing a dislike if they selected either the “dislike a great deal” or the “dislike somewhat” options for a particular genre.²

As we can see, Metal is the genre most likely to be disliked in both samples, consistent with previous work on Americans’ musical tastes based on earlier data (Bryson 1996; Lizardo and Skiles 2015). Both samples also agree that the two genres least likely to be disliked are Rock and Pop music. There is some disagreement across samples on the exact ranking of other genres. Country music experiences the most significant disparity, ranking second among the most disliked genres in the Prolific sample but near the bottom in the Lucid sample. Beyond that, both data sets agree that Musicals, Rap, and Disco are at the top of the dislike list, whereas Classical, R&B, and Jazz are closer to the bottom.

Univariate Distribution of Number of Dislikes

Table 2 shows the univariate distribution of the number of dislikes for the Prolific and Lucid samples. As we can see, both data sets closely agree on the proportion of

Table 1: Proportion of dislikes by genre.

Genre	Percentage Dislike (Prolific)	Genre	Percentage Dislike (Lucid)
Metal	38.4	Metal	42.6
Country	36.3	EDM	37.7
Musicals	33.9	Rap	36.2
Rap	31.2	Musicals	35.6
Disco	29.6	Disco	32.0
EDM	27.8	Reggae	28.5
Reggae	22.7	Jazz	25.8
Jazz	20.5	Classical	24.1
R_and_B	17.3	Country	22.1
Classical	13.9	R_and_B	17.3
Pop	9.7	Rock	14.3
Rock	5.3	Pop	12.9

Note:

N = 1408 (Prolific); N = 959 (Lucid); N = 2367 (All)

Table 2: Univariate distribution of dislikes for the Prolific and Lucid samples.

Count	Proportion (Prolific)	Proportion (Lucid)	Proportion (All)
0	0.192	0.201	0.196
1	0.126	0.117	0.122
2	0.167	0.132	0.153
3	0.156	0.141	0.150
4	0.135	0.111	0.125
5	0.097	0.095	0.096
6	0.055	0.066	0.060
7	0.033	0.047	0.038
8	0.018	0.028	0.022
9	0.011	0.028	0.018
10	0.010	0.016	0.012
11	0.000	0.013	0.005
12	0.001	0.006	0.003

Note:

N = 1408 (Prolific); N = 959 (Lucid); N = 2367 (All)

(potential) categorical tolerants, with the Prolific data yielding about 19 percent and the Lucid data set about 20 percent. When we combine both, we get an estimate of 19.6 percent, suggesting that about a fifth of the American population today may be classified as categorical tolerants, a presumed increase from L&S's estimate of about 15 percent using 2012 data.

As shown in Figure 1, the distribution of the number of dislikes we get from the pooled samples (like the separate sample distributions), exhibits excess zeroes, similar to the one obtained by L&S using 2012 data. Even considering issues of sample comparability relative to L&S's original study, we can conclude that the taste regime governing patterns of cultural choice among Americans in the second decade of the twenty-first century certainly no longer resembles that captured by the GSS in the early 1990s, with a larger-than-expected—if the number of dislikes followed a simple Poisson process—number of people refusing to express dislikes for any genre.

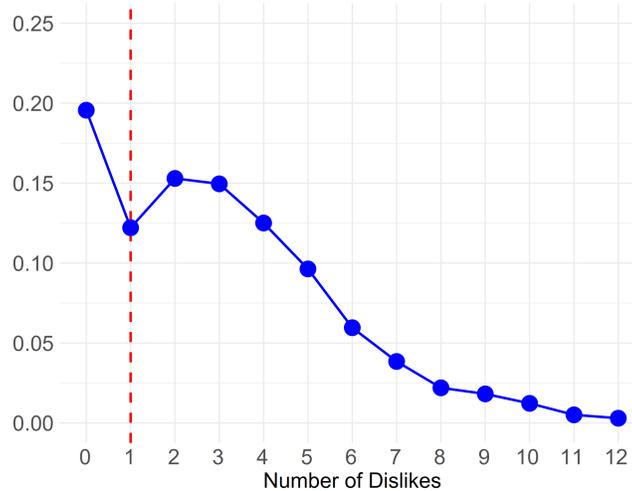


Figure 1: Univariate plot of number of dislikes, pooled Prolific and Lucid samples ($N = 2,367$).

However, one worry is that the differences between the percentages of categorical tolerants observed here and those in L&S’s original study are a simple function of the fact that the previous study used a larger number of genres than this one, thus exposing people to the “risk” of disliking more options, mechanically lowering the number of persons reporting zero dislikes in 2012 relative to the present study. As a robustness check, I compared the distribution of dislikes in the current sample with that in L&S’s original study.³ Although most of the abstract genre labels used in the two studies can be matched without a problem, two discrepancies need to be reconciled. First, the current research includes “Disco” as a genre, one that was absent from the original L&S study. Accordingly, I drop this genre, leaving us with 11 comparable genres. The other issue is that, while the current research asks only about likes and dislikes for a non-descript category of “Rock,” L&S’s previous study disaggregated this category into three subgenres: Classic, Contemporary, and Alternative. Accordingly, I merge these three categories in L&S’s original data into an omnibus “Rock” category, treating a dislike of any of them as a dislike of Rock. Finally, to compare apples to apples, I use only the more representative data collected in 2025 as a comparison point, as L&S’s data also came from a similar sample weighted to match 2010 census demographics in terms of age, race, and education.

With these caveats in mind, Figure 2 shows the univariate distribution in L&S’s original data ($N = 1,713$) against the 2025 Cint sample for the 11 matched genres. We can see that the estimated descriptive proportion of categorical tolerants is a bit higher than the number reported by L&S with the reduced number of genres (about 17 percent), suggesting that *some* of the observed differences may be due to this methodological artifact. However, note that the proportion of categorical tolerants is still significantly higher in the more recent data with the 11 matched genres (about 21 percent), with a Kolmogorov–Smirnov test rejecting the null hypothesis that the two samples are drawn from the same univariate distribution ($D = 0.09, p < 0.001$), suggesting that the distribution of dislikes has indeed shifted

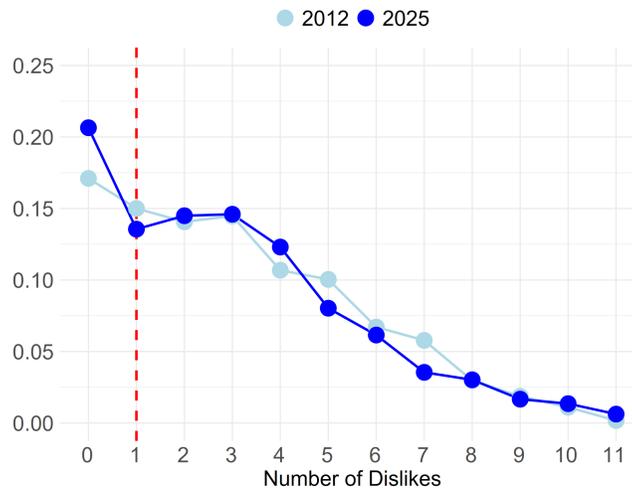


Figure 2: Univariate plot of number of dislikes, 2012 L&S sample (N = 1713) in light blue versus 2025 Cint Sample (N = 959) in blue.

since 2012. “Eyeballing” Figure 2 invites the inference that if there has indeed been such a shift, it is mainly concentrated in the estimated number of categorical tolerants.

Nevertheless, the *descriptive* proportion of persons with zero dislikes is not a strict estimate of the number of categorical tolerants, because these include a mixture of “structural zeros” (true categorical tolerants) and “random zeros” (people at risk of disliking but who did not). To tease this out and provide a more formal test of the differences between the 2012 and 2025 distributions of dislikes, I fit a zero-inflated Poisson model (see below for a more detailed justification of this statistical approach) to the pooled 2012/2025 data, with a single predictor in both the “count” and “zero-inflation” parts of the model: An indicator variable that equals one for observations in the 2025 Cint sample and zero for the ones in L&S’s 2012 sample. The results are shown in Figure 3a.

The single coefficient estimate in the “count” part of the model in Figure 3a tests the hypothesis that the differences between the two distributions are due to the non-zero counts (and whatever random zeros exist), a hypothesis not supported by the data ($\beta_{2025} = -0.01, p = 0.61$). The single coefficient estimate in the “inflation” part of the model tests the hypothesis that the main difference between the two distributions is the larger excess (structural) zeros observed in 2025, a hypothesis supported by the data at conventional levels of statistical significance ($\beta_{2025} = 0.26, p < 0.05$). Furthermore, the predicted probabilities from this model can give us a more precise estimate of the different representations of “true” categorical tolerants across the two samples; these are shown in Figure 3b.⁴ As we can see, the estimated number of categorical tolerants from the model in Table 3 for 2012 is close to that reported in L&S’s original article: about 15 percent. The number of categorical tolerants estimated in the 2025 Cint sample, however, is significantly larger, at just under 19 percent.

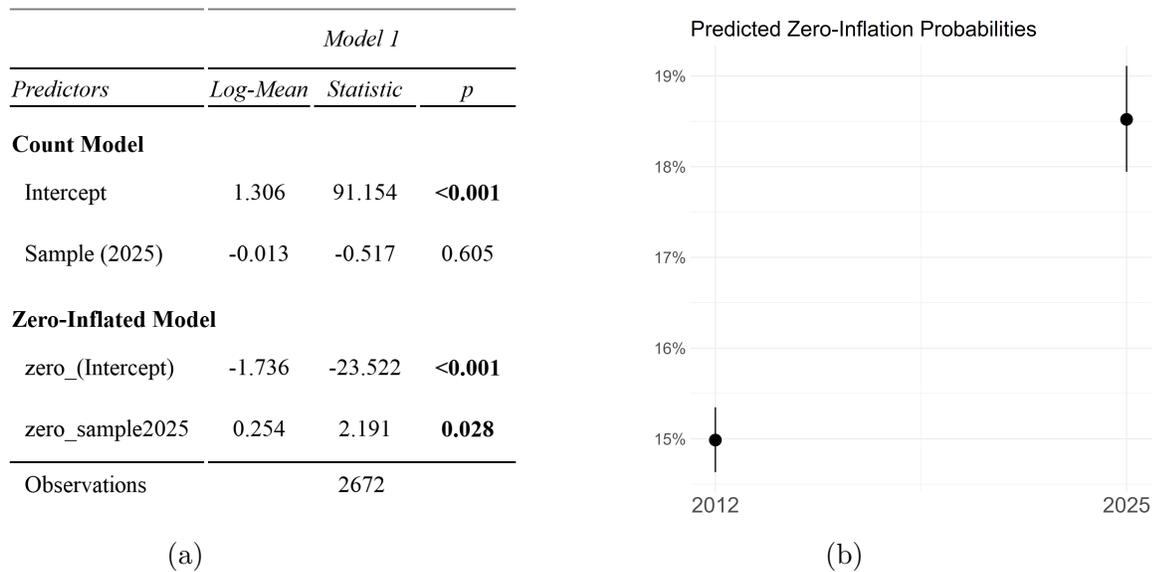


Figure 3: Coefficient estimate of sample membership comparing 2012 to 2025 from a zero-inflated Poisson model (a) and associated predicted probabilities of structural zeros for each sample (b).

Overall, the more conservative way of interpreting these results is that *some* of the differences between L&S's original estimates of the proportion of categorical tolerants in the U.S. population and the present study are due to the smaller number of genres included in the current survey. However, there is also sufficient evidence to suggest a significant (statistically and substantively) increase in the share of CTs over the ensuing decade and a half.

Socio-Demographic Distribution of CT

Figure 4 shows how the sample probability of being a categorical tolerant is distributed among various socio-demographic categories using the pooled ProA and Cint samples.⁵ Each panel of Figure 4 shows a "lollipop plot" with the numerical axis being the proportion of categorical tolerants at each level of the corresponding socio-demographic covariates. The red-dashed reference line shows the observed proportion of categorical tolerants in the sample.

As shown in Figure 4d, age is the most straightforward factor structuring CT, replicating the pattern observed by Lizardo and Skiles (2016b): Younger people are more likely to be categorical tolerants compared to older people, who are strongly under-represented among categorical tolerants; accordingly, the probability of being a categorical tolerant in these data is maximized among younger adults (individuals in their thirties), suggesting that very young people are induced to use musical genres for purposes of symbolic inclusion to a greater extent than people at this life-stage. Regardless, it is clear that if cohort replacement dynamics drive the rise of CT, we should expect it to remain a fixture of the U.S. cultural taste landscape as older, less categorically tolerant cohorts exit the population.

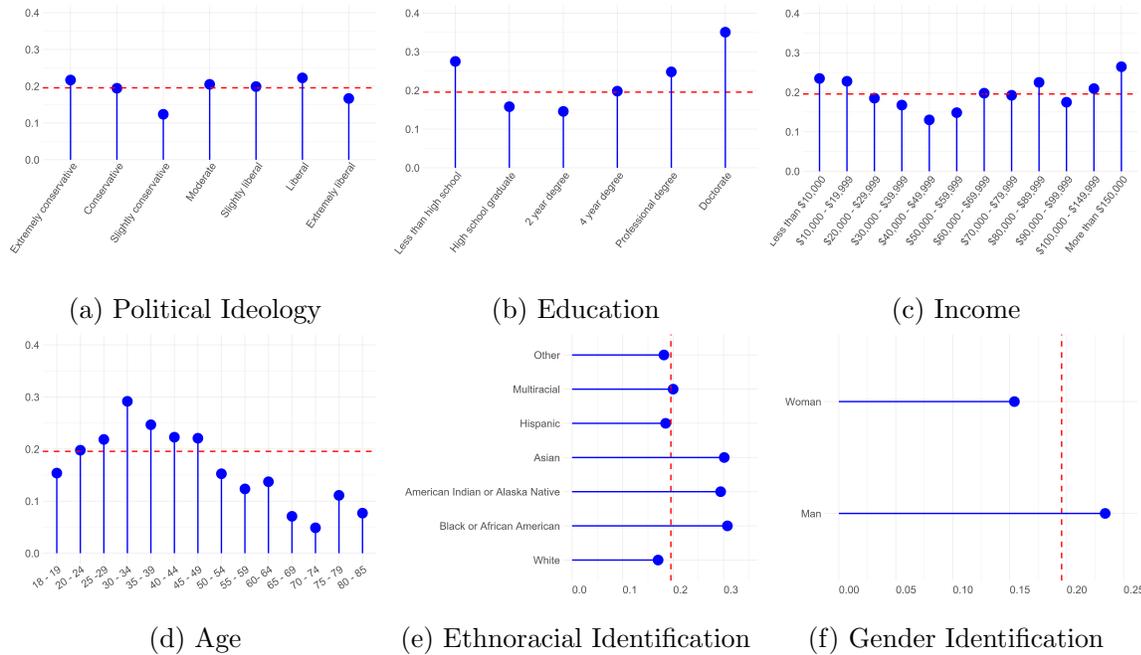


Figure 4: Proportion of categorical tolerant across different socio-demographic characteristics.

The other obvious factor structuring CT in the United States is race. Mirroring L&S's earlier findings, Figure 4e shows that CT continues to be primarily a non-white phenomenon, with whites (and to a similar extent respondents who identify as Hispanic) under-represented among categorical tolerant, and black, Asian, and Native American respondents over-represented. Also in line with L&S's findings, the traditional Bourdieusian structuring factors (e.g., education and income) do not have a straightforward link to CT, with both showing non-linear patterns (categorical tolerant over-represented among both high- and low-education and high- and low-income respondents).

Nevertheless, the apparent over-representation of the most highly educated stratum among the categorically tolerant suggests an elite move in this direction, which was not evident in L&S's previous analysis, which mainly pointed to cohort/age and ethnoracial status effects. Interestingly, we also observe a strong gender effect—unlike L&S's earlier findings—suggesting a relatively large gender gap in the CT taste pattern, with men overrepresented and women underrepresented among categorical tolerant. Finally, CT—in contrast to symbolic exclusion, as we will see in a bit—is not polarized along political-ideological lines, with no clear pattern by conservative/liberal placement.

Socio-Demographic Distribution of Symbolic Exclusion

The obverse phenomenon of CT is *symbolic exclusion* (Bryson 1996), that is, the probability of expressing at least one dislike, among those in the population who are at risk of doing so. Figure 5 shows the distribution of symbolic exclusion

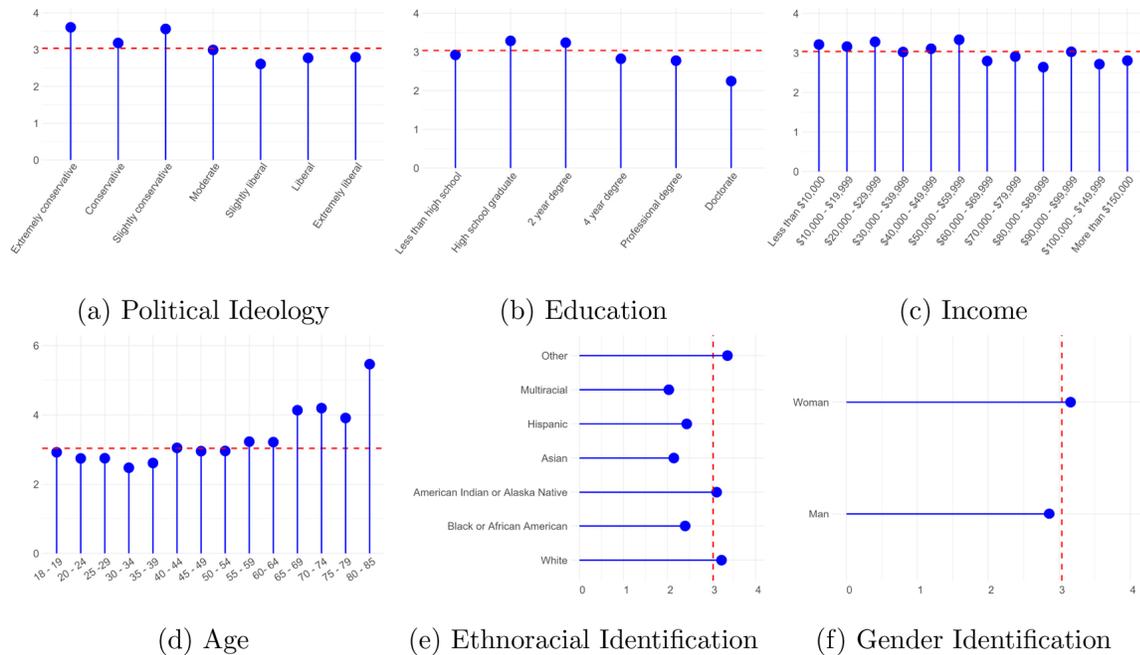


Figure 5: Average number of cultural dislikes across different socio-demographic characteristics.

propensities in the joint Prolific and Lucid samples using a similar plotting strategy as Figure 4, but this time with the average number of dislikes expressed by members of each socio-demographic category on the numerical axis. The dashed red reference line, like before, represents the average number of dislikes among individuals in the joint sample.

Symbolic exclusion tendencies across groups can be straightforwardly attributed to age and education, with older, less educated respondents expressing more dislike than younger, more educated ones. Importantly, symbolic exclusion is also structured by political ideology in the expected direction—see Rawlings and Childress (2023)—with liberals being less likely to express cultural dislikes than self-identified conservatives. In the same way, non-whites in the sample are less likely to express cultural dislikes (compared to white respondents), whereas women express more dislikes (compared to those who identify as men).

Multivariable Analysis Results

In their original article, Lizardo and Skiles (2016b) proposed modeling CT and symbolic exclusion jointly as a hurdle process. The basic idea is that categorical tolerants are a distinct slice of the population who are not at risk of engaging in symbolic exclusion (expressing dislikes). Thus, we can statistically model the joint predictors of both tendencies by assuming that there is a set of predictors that drive the probability of being in the categorical tolerant group (thus not being at risk of expressing dislikes), and net of this, there is a second set of predictors that drive the number of dislikes expressed for those who are at risk of doing so. The zero-inflated

Poisson model is thus a natural statistical framework for modeling this dual regime data-generating process (Zorn 1998). Because the univariate distribution of cultural dislikes is comparable across samples (see 2), I pool the three samples to increase statistical power and report results from the pooled sample analysis, including a binary indicator for each survey sample (using the 2021 Prolific sample as the reference category) in each model. I use listwise deletion of cases with missing values for the socio-demographic covariates included in the model, which leaves us with a total analytical sample of $N = 2,291$.

In line with this analytic strategy, Table 3 presents the coefficient estimates from four zero-inflated Poisson regression models. The specifications in each model, for both the count and zero-inflated components, are informed by the descriptive patterns we considered in Figures 4 and 5. Coefficient estimates corresponding to the count part of the model—predicting the number of dislikes among those who are at risk of expressing them—are shown in the top half of the table. Coefficient estimates for the zero-inflation part of the model, predicting the odds of landing in the categorical tolerant group—and thus not being at risk of engaging in symbolic exclusion—are presented in the lower half of the table. In all models, age, education, income, and political ideology (with higher values indicating a more liberal stance) are entered as ordered numerical inputs. Race (white) and gender (Women) are entered into the model as dummy indicators. The zero-inflated specifications also contain squared terms for education and income to model expected curvilinear effects (see Figures 4b and 4c).

Let us begin with the predictors of symbolic exclusion. Model 1 includes all of the main socio-demographic predictors. The results align with expectations derived from the earlier descriptive patterns. Age increases the likelihood of engaging in symbolic exclusion, but education and income reduce this tendency. Nevertheless, the education effect, as suggested by Figure 4b, is substantively small and not statistically significant ($p = 0.09$) using the traditional cutoff value of $p < 0.05$, and becomes smaller and crosses the threshold to decidedly not statistically significant in later specifications. Being politically liberal, on the other hand, has a substantively significant effect, one that is robust—in terms of statistical significance—across models. As suggested by Figure 4a, self-identified Liberals are less likely to engage in symbolic exclusion than conservatives. These results indicate that in the United States today, political liberalism is a more reliable predictor of whether a person will engage in symbolic exclusion than is educational attainment, a situation quite different from that which obtained in the 1990s (Bryson 1996). When it comes to ethnoracial identity, self-identified white respondents are more likely to express dislike than respondents who identified with one of the non-white ethnoracial categories ($p < 0.01$). Finally, there is no appreciable gender-identity difference in the propensity to engage in symbolic exclusion via cultural dislikes ($p = 0.53$).

Model 1 includes age as the sole demographic predictor in the zero-inflated part of the model, providing strong evidence that CT is primarily a cultural habit among the young and is less prevalent among older respondents ($p < 0.001$). Model 2 includes race and gender as additional predictors of CT, verifying the expectation that white respondents are less likely to be found in this group (compared to non-white respondents) and women are less likely to be categorical tolerants (compared

Table 3: Coefficient estimates from zero-inflated Poisson models predicting CT and symbolic exclusion, joint prolific and Lucid samples.

<i>Predictors</i>	<i>Model 1</i>		<i>Model 2</i>		<i>Model 3</i>		<i>Model 4</i>	
	<i>Log-Mean</i>	<i>p</i>	<i>Log-Mean</i>	<i>p</i>	<i>Log-Mean</i>	<i>p</i>	<i>Log-Mean</i>	<i>p</i>
Intercept	1.247	<0.001	1.266	<0.001	1.250	<0.001	1.252	<0.001
Age	0.014	0.001	0.015	0.001	0.014	0.001	0.014	0.001
Education	-0.020	0.090	-0.019	0.102	-0.012	0.313	-0.012	0.321
Income	-0.012	0.004	-0.012	0.004	-0.011	0.005	-0.012	0.003
Pol. Liberalism	-0.035	<0.001	-0.036	<0.001	-0.036	<0.001	-0.036	<0.001
White	0.201	<0.001	0.184	<0.001	0.184	<0.001	0.185	<0.001
Woman	0.016	0.530						
Survey (Prolific 2023)	0.072	0.045	0.071	0.047	0.065	0.067	0.066	0.065
Survey (Lucid 2025)	0.116	0.002	0.113	0.003	0.114	0.002	0.114	0.002
Zero-Inflated Model								
zero_Intercept	-0.795	<0.001	-0.419	0.015	-0.110	0.827	0.243	0.640
zero_Age	-0.114	<0.001	-0.097	<0.001	-0.118	<0.001	-0.117	<0.001
zero_Survey (Prolific 2023)	-0.505	0.003	-0.485	0.004	-0.490	0.004	-0.482	0.005
zero_Survey (Lucid 2025)	0.057	0.713	0.057	0.716	0.324	0.052	0.320	0.057
zero_White			-0.327	0.021	-0.354	0.013	-0.329	0.023
zero_Woman			-0.553	<0.001	-0.502	<0.001	-0.512	<0.001
zero_Education					-0.561	0.051	-0.396	0.178
zero_Education Squared					0.119	0.003	0.099	0.013
zero_Income							-0.236	0.004
zero_Income Squared							0.016	0.009
Observations	2291		2291		2291		2291	
R ² / R ² adjusted	0.218 / 0.215		0.237 / 0.234		0.263 / 0.260		0.264 / 0.261	

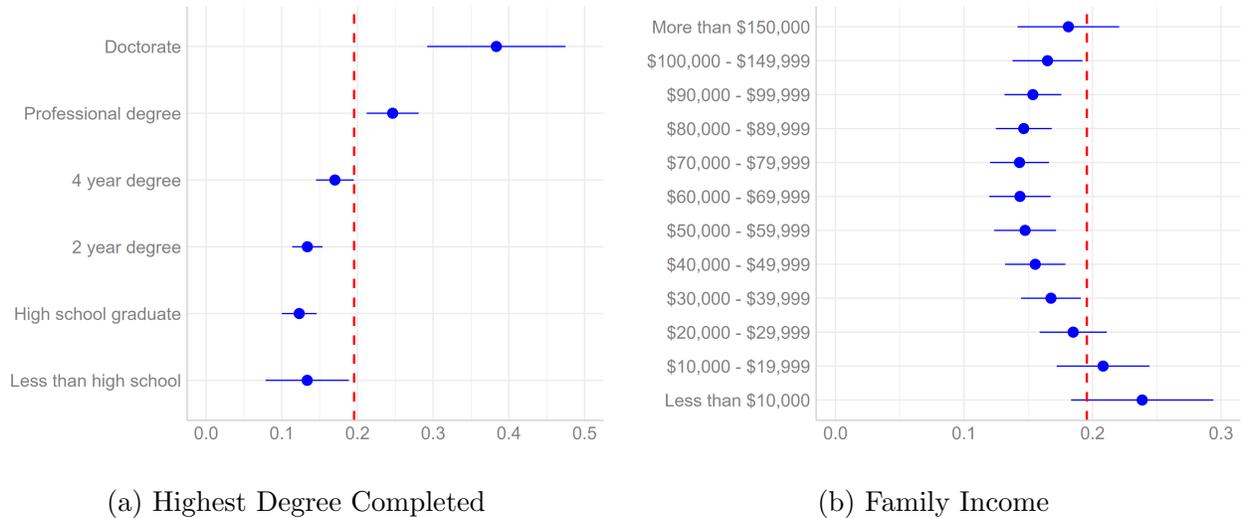


Figure 6: Predicted marginal means of the probability of being a categorical tolerant for education and income groups obtained from the estimates of model 4 of Table 3.

to respondents who identify as men). Model 3 introduces education as an additional predictor of CT, including a linear input for the respondent's education and a square term for the same covariate. The results verify the idea, suggested by Figure 4b, the education effect on the odds of being a categorical tolerant is parabolic, with a negative linear component (indicating that lower education person are more likely to be in the categorical tolerant group) and a positive squared component (indicating that at some point the negative effect reverses, and very high-education individuals end up over-represented among the categorical tolerants). Model 4 adjusts for the non-linear effect of self-reported family income on the odds of being a categorical tolerant. Both components of the income covariate have a statistically significant impact on the odds of being classified as a CT, with a negative linear effect and a positive quadratic effect, suggesting an under-representation of people in middle-income categories among categorically tolerant individuals.

To facilitate the interpretation of the non-linear education and income effects, Figure 6 shows marginal effect estimates for the probability of ending up in the categorically tolerant group, holding all other predictors at their observed means.⁶ As shown in Figure 6a, while the effect of education is non-linear, it is monotonic, suggesting that people with the most advanced levels of education (Professional degree or PhD) are over-represented among the categorical tolerants. This contrasts with the pattern of results obtained by L&S using 2012 data, suggesting a Bourdieusian upgrading of CT as an elite taste pattern in the United States today. On the other hand, the income effect is parabolic along the observed range. This suggests that middle-income people are underrepresented among the categorically tolerant, without implying any over-representation of very low or very high-income people in this category, who are instead at the observed average proportion for the sample.

Discussion and Concluding Remarks

Summary of Key Results

The findings presented in this article offer a critical update to recent work on changing cultural taste patterns in the United States, particularly regarding the dynamic processes of inclusion and exclusion in cultural preferences centered around the idea of CT. The core empirical results further establish the enduring relevance and continuous diffusion of CT, a cultural pattern first identified in earlier work by L&S (2016b). Specifically, approximately 20 percent of respondents across samples collected between 2021 and 2025 refused to report disliking any of the measured cultural genres, a larger proportion than observed in previous work by L&S. This basic descriptive result confirms that the univariate distribution of dislikes maintains the expected “zero-inflated” structure, diverging from a simple Poisson process, as observed earlier for the 1993 GSS data.

The persistence of CT in the United States is strongly linked to cohort replacement dynamics, as evidenced by the over-representation of younger adults among categorical tolerants relative to older cohorts. Furthermore, consistent with previous analyses, non-white respondents remain overrepresented within the CT group. This tendency suggests that the refusal to engage in symbolic exclusion persists as a macro-level pattern of cultural change, driven by younger non-white respondents. A particularly salient development in the social patterning of CT in the current study is its observed “trickle-up” into higher-educated strata, indicating potentially shifting bases of elite cultural engagement. Unlike L&S’s earlier findings, which noted a slight under-representation of college-educated individuals among CTs, the current results show that those reporting the highest levels of education are now *over-represented* in this group. This implies that Bourdieusian dynamics of status display, characterized by an embrace of “conspicuous openness to diversity,” may be increasingly at play, positioning CT as a preferred cultural tool for contemporary elites (Jarness and Friedman 2017; Ollivier 2008).

Beyond status and demographic correlates, the study offers insight into the growing linkage between cultural taste and political orientation (Rawlings and Childress 2023). Although CT itself does not appear to be polarized along liberal/conservative lines, the propensity for symbolic exclusion—the act of expressing dislikes—has evidently become polarized among the subset of the population at risk of expressing dislikes. After adjusting for other sociodemographic factors, conservatives are more likely to express a greater number of cultural dislikes than liberals, with the political ideology effect dwarfing the effect of education—the most reliable predictor of this phenomenon in previous work (Bryson 1996)—in both substantive and statistical significance. This suggests that political ideology, alongside age, has become a reliable predictor of dislike for cultural genres in the contemporary United States, echoing recent work that identifies the increasing role of political polarization in shaping cultural taste patterns in the United States (Rawlings and Childress 2023).

Study Limitations

One limitation of the present study is the need to compare findings over time using survey instruments with varying scopes. The present study's cultural taste measures rely on fewer genres (12) than L&S's original research (18 in the 2012 survey). This reduction in the item battery invites caution when making direct comparisons of dislike volumes across time. That said, the robustness check provides reassuring evidence for the inference of a discernible secular increase in the proportion of the population classifiable as CT using genres matched to those considered in the current study, relative to L&S's earlier work.

The theoretical interpretation of CT's meaning—its Simmelian “content” as compared to its “form” (Sonnnett 2004)—remains an area of ongoing sociological investigation and debate, complicated by inherent data constraints.⁷ Some comparative research suggests that the unwillingness to dislike may sometimes indicate non-committal behaviors such as indecisiveness or generalized ambivalence (often captured as “mixed feelings” or “do not agree nor disagree”), rather than authentic cosmopolitan tolerance or a genuine embrace of cultural diversity. Indeed, the absence of dislikes may signal a socially unstructured ambivalence that increased over time in other contexts. I used the zero-inflated Poisson model to statistically separate the population at risk of exclusion from the categorical tolerants, showing that the latter occupy a coherent social niche. Nevertheless, the underlying disposition and discursive positioning—whether strategic status signaling or genuine universal acceptance—cannot be definitively determined using purely survey-based methods; instead, qualitative approaches such as in-depth interviews are better suited to extracting the meaning behind more abstract patterns of cultural choice (Atkinson 2011; Jarness and Friedman 2017; Lamont and Swidler 2014).

In addition, while broad genre categories are standard in large-scale research on the sociology of taste, this approach necessarily obscures nuances critical to contemporary distinction. Critics point out that focusing solely on genre acceptance risks overlooking within-genre hierarchies, where specific tastes for objects (artists, tracks, and specific works) within accepted broad genres are increasingly crucial for symbolic exclusion and status signaling among higher-status groups (Childress and Nault 2019; Nault et al. 2021). Using broad genre categories may therefore limit the ability to fully capture the subtle, exclusive tastes that coexist with inclusive genre preferences in the contemporary configuration of elite cultural consumption (Lizardo 2024).

Suggestions for Future Work

The current research suggests several promising directions for future research that aim to deepen our theoretical and empirical understanding of taste, tolerance, and inequality in the contemporary moment. First, moving forward requires adopting more fine-grained measurement strategies to overcome the limitations of broad genres. Future research should replicate this work using more fine-grained genre categories, perhaps focusing on objects (artists, albums, and specific works) rather than just genres, to comprehensively examine the hypothesized contemporary configuration of high-status tastes. Is CT only observed among broad genre

categories, or can it also be observed at the level of objects within genres? Such refined analyses are increasingly relevant for understanding how cultural boundaries persist or are redrawn in cultural fields.

Second, more targeted research, using interview-based or focus group strategies, is required to definitively unpack the sociological meaning of CT. Future studies should move beyond measuring the mere volume of likes and dislikes toward qualitative or mixed-method approaches that capture the underlying motivations or the specific “mode” of appropriation of CTs relative to particular objects and practices. This would distinguish between CT as a performance of socially valued openness (e.g., a strategic status performance), a broader cosmopolitan disposition, or simply ambivalence/indecisiveness in taking a stance.

Third, given the reliable connection between symbolic exclusion and political ideology, further scholarly attention must be dedicated to tracking this accelerated linkage. Research should continue to monitor how cultural taste patterns are drawn into the “oil spill” of political polarization (DellaPosta 2020). This could involve investigating the processes and mechanisms via which political actors may strategically impose cultural boundaries, thereby reinforcing political cleavages via cultural affinity or distancing.

Finally, future sociological inquiry should broaden its scope to include comparative and contextual replication. Although the U.S. context provides specific insights into racial dynamics and high polarization, examining how CT manifests in other nations—such as the recent comparative framework in Finland or social settings outside the Euro-American West like South Korea—could illuminate how local social topography and other divisions modulate the social correlates and meaning of CT. Furthermore, exploring whether personality traits, such as openness or extroversion (Andrews et al. 2022), mediate the relationship between cultural exposure and the resulting omnivorous or categorically tolerant dispositions could offer a promising interdisciplinary line of inquiry.

Notes

- 1 This last is what Vries and Reeves (2022) call the “weak” omnivore thesis, but which like many things weak in sociology, it is also the only empirically plausible version of the thesis.
- 2 The genres included in previous work that are not included in the present study are Big Band/Swing, Bluegrass, Folk, Gospel, Latin/Mariachi/Salsa, Mood/Easy Listening, and New Age/Space Music.
- 3 I thank an anonymous reader for suggesting this approach.
- 4 Estimated using the `ggeffects` package in *R*.
- 5 As noted, the descriptive proportion of zeros in the sample is an approximation of the quantity we seek, but still helpful to determine key trends. The multivariable results from the zip model in Table 3 provide the more definitive test.
- 6 Estimated using the `ggeffects` package in *R*.
- 7 Another option is that the refusal to dislike is itself a Simmelian “contentful form” making an independent semiotic statement aside from content (Lizardo 2019).

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